

# Portfolio Optimization with Stochastic Volatilities : A Backward Approach

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# Introduction

Portfolio optimization: a fundamental concern when investors trade between a large number of risky assets

- Merton (1971): Maximizing expected utility of terminal wealth  
He derived a closed formula.
- Many generalizations: maximizing under Bankruptcy constraint, maximizing with transaction costs, maximizing in pure-jump Lévy model.

**Maximizing utility problems** lead to **HJB** Equations

An efficient algorithm: **Howard algorithm** computing two sequences the optimal strategy and the value function. **Not possible when risky assets number  $> 3$ .**

# Introduction

- Zariphopoulou (2001): maximizing utility in a market with one risky asset and a stochastic volatility model. Using a power transformation: the value function is a solution of a linear P.D.E
- Pham (2002): multidimensional model, diffusion process, constraints. A power transformation leads to a **semi linear parabolic equation**.
- Mnif ((2007): multidimensional model, stochastic volatility model, exponential utility, constraints on amounts, jump diffusion process. When there is no correlation between the risky asset and the factor volatility, the optimal investment strategy is determined by solving a static optimization problem.

# Problem formulation

$(\Omega, \mathcal{F}, \mathcal{P})$  filtered probability space

financial market: bond  $S^0$  and  $n$  risky assets  $S$

$$\begin{aligned}
 S^0 &\equiv 1 \\
 dS_t &= \text{diag}(S_{t-}) \left( b(\Lambda_t)dt + \sigma(\Lambda_t)dW_t + \bar{\sigma}(\Lambda_t)d\bar{W}_t \right. \\
 &\quad \left. + \int_{R^n \setminus \{0\}} \gamma(\Lambda_t, u) \tilde{\mu}(dt, du) \right)
 \end{aligned}$$

$W$   $d$ -dimensional standard Brownian motion

$\bar{W}$   $m$ -dimensional standard Brownian motion independent of  $W$

$\mu$  a Poisson random measure and

$\tilde{\mu}(dt, du) = (\tilde{\mu}(dt, du))_{1 \leq i \leq n} = (\mu_i(dt, du) - q_i(du)dt)_{1 \leq i \leq n}$  is

the compensated Poisson random measure

# Problem formulation

$q_i(du)$  is the Lévy measure

$$\int_{\mathbb{R}^n \setminus \{0\}} q_i(du) < \infty.$$

$\Lambda$  is a  $d$  dimensional stochastic factor

$$d\Lambda_t = \eta(\Lambda_t)dt + dW_t,$$

Assumption **(H1)**  $\eta$  is Lipschitz.

We denote

$$\Sigma(\lambda) = (\sigma(\lambda), \bar{\sigma}(\lambda)),$$

$$\alpha(\lambda) = \inf_{\pi \in \mathbb{R}^n, \pi \neq 0} \frac{|\Sigma(\lambda)^* \pi|^2}{|\pi|^2}, \quad \lambda \in \mathbb{R}^d,$$

# Problem formulation

Assumption **(H2)** There exists some positive constants  $C$  such that for a.e.  $(\lambda, u) \in R^d \times R^n \setminus \{0\}$ :

$$\inf_{\lambda \in R^d} \alpha(\lambda) > 0$$

$$|b(\lambda)| + \|\sigma(\lambda)\| \leq C.$$

$$\|\bar{\sigma}(\lambda)\| \leq C(1 + |\lambda|)$$

$$\gamma_{ij}(\lambda, u) > -1 \text{ for all } 1 \leq i, j \leq n$$

$$\|\gamma(\lambda, u)\| \leq C$$

# Problem formulation

A strategy  $\pi$  is said admissible if it is  $F$ -predictable and

$$\pi_t \in \Delta := \{(\pi_1, \dots, \pi_n) \in [0, 1]^n, \sum_{i=1}^n \pi_i \leq 1\} \text{ a.s. for all } 0 \leq t \leq T.$$

$\mathcal{A}$  the set of admissible controls.

$$\begin{aligned} X_t = & x + \int_0^t X_{s-} \left( \pi_s^* b(\Lambda_s) ds + \pi_s^* \sigma(\Lambda_s) dW_s + \pi_s^* \bar{\sigma}(\Lambda_s) d\bar{W}_s \right. \\ & \left. + \int_{\mathbb{R}^n \setminus \{0\}} \pi_s^* \gamma(\Lambda_s, u) \tilde{\mu}(ds, du) \right) \end{aligned}$$

The utility function:  $U(x) = \frac{x^\delta}{\delta}$ ,  $x \in \mathbb{R}_+$ , The objective of the agent

$$v(t, x, \lambda) = \sup_{\pi \in \mathcal{A}} E[U(X_T) | X_t = x, \Lambda_t = \lambda], \quad (t, x, \lambda) \in [0, T] \times \mathbb{R}_+ \times \mathbb{R}^d.$$

# Semi-linear PDE

$$\begin{aligned}
 & \frac{\partial v}{\partial t} + \eta(\lambda)^* D_\lambda v + \frac{1}{2} \Delta_\lambda v + \max_{\pi \in \Delta} \left\{ x \pi^* b(\lambda) \frac{\partial v}{\partial x} \right. \\
 & + x^2 \frac{1}{2} |\Sigma(\lambda)^* \pi|^2 \frac{\partial^2 v}{\partial x^2} + x \pi^* \sigma(\lambda) D_{x\lambda}^2 v \\
 & + \sum_{i=1}^n \int_{R^n \setminus \{0\}} (v(t, x(1 + \pi^* \gamma_i(\lambda, u)), \lambda) \\
 & \left. - v(t, x, \lambda) - x \pi^* \gamma_i(\lambda, u) \frac{\partial v}{\partial x} \right) q_i(du) \Big\} = 0,
 \end{aligned}$$

for a.e.  $(t, x, \lambda) \in [0, T) \times R_+ \times R^d$

$$v(T, x, \lambda) = \frac{x^\delta}{\delta},$$

# Semi-linear PDE

Due to the power utility function, we look for a candidate of HJB equation in the form:

$$v(t, x, \lambda) = \frac{x^\delta}{\delta} \exp(-\phi(t, \lambda)).$$

By differentiation, HJB implies

$$-\frac{\partial \phi}{\partial t} - \frac{1}{2} \Delta \phi + H(\lambda, D\phi) = 0, \quad (t, \lambda) \in [0, T] \times \mathbb{R}^d, \quad (1)$$

with terminal condition

$$\phi(T, \lambda) = 0, \quad y \in \mathbb{R}^d, \quad (2)$$

# Semi-linear PDE

$H$  is defined on  $R^d \times R^d$  by

$$\begin{aligned}
 H(\lambda, p) &= \frac{1}{2}|p|^2 - p^* \eta(\lambda) \\
 &+ \max_{\pi \in \Delta} \left\{ \delta(\pi^* b(\lambda) - \pi^* \sigma(\lambda) p) - \frac{1}{2} \delta(1 - \delta) |\Sigma(\lambda)^* \pi|^2 \right. \\
 &+ \left. \sum_{i=1}^n \int_{R^n \setminus \{0\}} \left( (1 + \pi^* \gamma_i(\lambda, u))^\delta - 1 - \delta \pi^* \gamma_i(\lambda, u) \right) q_i(du) \right\}.
 \end{aligned}$$

# Semi-linear PDE

## Theorem

*Let assumptions **(H1)** and **(H2)** hold. Suppose that there exists a solution  $\phi \in C^{1,2}([0, T] \times R^d) \cap C^0([0, T] \times R^d)$  to the semi-linear equation (1) with the terminal condition (2). We also assume that  $|D\phi(t, \lambda)|$  has a linear growth condition in  $\lambda$  uniformly in  $t$ . Then the value function of (1) is given by:*

$$v(t, x, \lambda) = \frac{x^\delta}{\delta} \exp(-\phi(t, \lambda)), \quad (t, x, \lambda) \in [0, T] \times R \times R^d.$$

## Theorem

The optimal portfolio is given by the Markov control  
 $\{\hat{\pi}_t = \hat{\pi}(t, \Lambda_t), 0 \leq t \leq T\}$  where

$$\hat{\pi}(t, \lambda) \in \left. \begin{aligned} & \arg \min_{\pi \in \Delta} \left\{ -\delta(\pi^* b(\lambda) - \pi^* \sigma(\lambda) D\phi(t, \lambda)) \right. \\ & + \frac{1}{2} \delta(1 - \delta) |\Sigma(\lambda)^* \pi|^2 \\ & \left. - \sum_{i=1}^n \int_{R^n \setminus \{0\}} \left( (1 + \pi^* \gamma_i(\lambda, u))^\delta - 1 - \delta \pi^* \gamma_i(\lambda, u) \right) q_i(du) \right\}. \end{aligned} \right\}$$

# Regularity of the value function

**(H3)**(i)  $\inf_{\lambda \in R^d} \alpha(\lambda) > 0$  (uniform elliptic volatility).

(ii)  $\eta$  is  $C^1$  and Lipschitz,  $b$  is  $C^1$  and bounded and  $Db$  is bounded.

(iii)  $\Sigma\Sigma^*$  is  $C^1$  and bounded and  $\|D(\Sigma\Sigma^*)\|$  is bounded

(iv)  $-1 < \gamma_{ij}(\lambda, u) = \gamma_{ij}(u) \leq M$  for all  $(\lambda, u) \in R^d \times R^n \setminus \{0\}$  where  $M > 0$ .

# Regularity of the value function

## Theorem

*Under Assumption **(H3)**, there exists a solution  $\phi \in C^{1,2}([0, T] \times R^d) \cap C^0([0, T] \times R^d)$  to the semi-linear equation (1) with terminal condition (2) and linear growth condition in  $\lambda$  uniformly in  $t$  on the derivative  $D\phi$ .*

$$\begin{aligned} \frac{dS_t^i}{S_{t-}^i} &= b_i dt + \nu_i(\Lambda_{it}) \sum_{j=1}^i \rho_{ij} (\rho_j dW_t^j + \sqrt{1 - \rho_j^2} d\bar{W}_t^j) \\ &+ \sum_{j=1}^n \int_{\mathbb{R}^n \setminus \{0\}} \gamma_{ij}(u) \tilde{\mu}_j(dt, du), \\ d\Lambda_{it} &= (a_i - \theta_i \Lambda_{it}) dt + dW_t^i, \text{ for all } i \in \{1, \dots, n\}. \end{aligned}$$

$(\nu_i)_{1 \leq i \leq n}$  bounded  $C^1$  functions with bounded derivatives and lower bounded by a positive constant  $\epsilon_i > 0$  for all  $i \in \{1, \dots, n\}$   
 $a_i$  and  $\theta_i$  are constants,  $\rho_{ij}$  is the constant correlation between the two Brownian motions of  $S^i$  and  $S^j$ ,  $\rho_i$  is the constant correlation between  $S^i$  and its volatility and  $\gamma_{ij}(u) > -1$  and is bounded.

The optimal portfolio is given by:

$$\begin{aligned}
 \hat{\pi}(t, \Lambda_t) \in \arg \min_{\pi \in \Delta} & \left[ -\delta \sum_{i=1}^n \pi_i b_i \right. \\
 + & \frac{\delta(1-\delta)}{2} \sum_{i,j=1}^n \pi_i \pi_j \left( (\nu_i(\Lambda_{it})) (\nu_j(\Lambda_{jt})) \sum_{k=1}^{\inf(i,j)} \rho_{ik} \rho_{jk} \right) \\
 + & \delta \sum_{i=1}^n \sum_{j=1}^i \pi_i \rho_{ij} \rho_j (\nu_i(\Lambda_{it})) \frac{\partial \phi}{\partial \lambda_j}(t, \Lambda_t) \\
 - & \left. \sum_{i=1}^n \int_{R^n \setminus \{0\}} \left( (1 + \pi^* \gamma_i(u))^\delta - 1 - \delta \pi^* \gamma_i(u) \right) q_i(du) \right].
 \end{aligned}$$

We define two processes  $(Y, Z)$  by

$$Y_t := \phi(t, \Lambda_t), \quad Z_t := D\phi(t, \Lambda_t) \text{ for all } t \in [0, T].$$

### Proposition

*Suppose that Assumptions **(H1)**, **(H2)** and **(H3)** hold. We define the processes  $(\bar{Y}, \bar{Z})$  as follows*

$$\bar{Y}_t := \exp(Y_t), \quad \bar{Z}_t := \bar{Y}_t Z_t, \text{ for all } t \in [0, T],$$

*then the couple  $(\bar{Y}, \bar{Z})$  is a solution of the following BSDE*

$$-d\bar{Y}_t = \bar{g}(\Lambda_t, \bar{Y}_t, \bar{Z}_t)dt - \bar{Z}_t dW_t,$$

*with terminal condition*

## Proposition

$$\bar{Y}_T = 1$$

where

$$\begin{aligned} \bar{g}(\lambda, \bar{y}, \bar{z}) = & - \min_{\pi \in \Delta} \left\{ -\delta(\pi^* b(\lambda) \bar{y} - \pi^* \sigma(\lambda) \bar{z}) \right. \\ & + \bar{y} \frac{\delta(1-\delta)}{2} |\Sigma(\lambda)^* \pi|^2 \\ & \left. - \bar{y} \sum_{i=1}^n \int_{R^n \setminus \{0\}} \left( (1 + \pi^* \gamma_i(u))^\delta - 1 - \delta \pi^* \gamma_i(u) \right) q_i(du) \right\}. \end{aligned}$$

# Discretization and simulation of the decoupled FBSDE

**Step 1. Problem discretization.**  $(t_k := kh = \frac{kT}{N})_{0 \leq k \leq N}$ .

$$\begin{aligned}\Lambda_{l,t_0}^N &= \Lambda_{l,0} \\ \Lambda_{l,t_{k+1}}^N &= \Lambda_{l,t_k}^N + \eta(\Lambda_{l,t_k}^N)h + \Delta W_{l,k},\end{aligned}$$

$$\Delta W_{l,k} = W_{l,t_{k+1}} - W_{l,t_k}, \quad 1 \leq l \leq d.$$

$$\begin{aligned}\bar{Y}_{t_N}^N &= 1 \\ \bar{Z}_{l,t_k}^N &= \frac{1}{h} E_k(\bar{Y}_{t_{k+1}}^N \Delta W_{l,k}), \\ \bar{Y}_{t_k}^N &= E_k(\bar{Y}_{t_{k+1}}^N) + h E_k(\bar{g}(\Lambda_{t_k}^N, \bar{Y}_{t_{k+1}}^N, \bar{Z}_{t_k}^N)),\end{aligned}$$

where  $E_k(\cdot) = E(\cdot | \mathcal{F}_{t_k})$ .

## Error induced the discretization in time

### Remark

*Since  $Y_{t_N}^N$  is a constant and  $\Lambda^N$  is a Markov Chain, it is easy to see that  $\bar{Y}_{t_k}^N = \bar{y}_k^N(\Lambda_{t_k}^N)$  and  $\bar{Z}_{t_k}^N = \bar{z}_k^N(\Lambda_{t_k}^N)$  where  $\bar{y}_k^N$  and  $\bar{z}_k^N$  are unknown regression functions defined in a backward manner.*

Zhang (2004), Bouchard and Touzi (2004), Gobet et al. (2005) proved For  $h$  small enough

### Theorem

$$\max_{0 \leq k \leq N} E |\bar{Y}_{t_k} - \bar{Y}_{t_k}^N|^2 + \sum_{k=0}^{N-1} \int_{t_k}^{t_{k+1}} E |\bar{Z}_t - \bar{Z}_{t_k}^N|^2 dt \leq C (1 + |\Lambda_0|^2) h.$$

# Discretization and simulation of the decoupled FBSDE

**Step 2. Localization.** Localization of the Brownian increments and the function  $\bar{g}$

- $[\Delta W_{l,k}]_w = -R_0 \sqrt{(h)} \vee \Delta W_{l,k} \wedge R_0 \sqrt{(h)}$
- $\bar{g}^R(\lambda, \bar{y}, \bar{z}) = \bar{g}^R(-R_1 \vee \lambda_1 \wedge R_1, \dots, -R_d \vee \lambda_d \wedge R_d, \bar{y}, \bar{z})$

The localization modifies the numerical scheme. We define  $(\bar{Y}^{N,R}, \bar{Z}^{N,R})$  by

$$\bar{Y}_{t_N}^{N,R} = 1$$

$$\bar{Z}_{l,t_k}^{N,R} = \frac{1}{h} E_k(\bar{Y}_{t_{k+1}}^{N,R} [\Delta W_{l,k}]_w),$$

$$\bar{Y}_{t_k}^{N,R} = E_k(\bar{Y}_{t_{k+1}}^{N,R}) + h E_k(\bar{g}^R(\Lambda_{t_k}^N, \bar{Y}_{t_{k+1}}^{N,R}, \bar{Z}_{t_k}^{N,R})),$$

# Discretization and simulation of the decoupled FBSDE

## Remark

*The main interest of the localization is to provide bounded regression functions  $\bar{y}_k^{N,R}$  and  $\bar{z}_k^{N,R}$ . One has*

*$\|\bar{y}_k^{N,R}\|_\infty \leq C_y(R)$  and  $\|\bar{z}_k^{N,R}\|_\infty \leq C_z(R)$  ( See Proposition 1 (Lemor Gobet and Warin (2006))*

## Step 3. Function bases.

- We approximate  $\bar{Y}_{t_k}^{N,R}$  and  $\bar{Z}_{l,t_k}^{N,R}$  for all  $1 \leq l \leq d$  by a projection on a finite-dimensional function bases  $p_{0,k}(\Lambda_{t_k}^{N,R})$  and  $p_{l,k}(\Lambda_{t_k}^{N,R})$  for all  $1 \leq l \leq d$ .
- We set  $\alpha_{l,k}$  for all  $0 \leq l \leq d$  the projection coefficients of  $\bar{Y}_{t_k}^{N,R}$  and  $\bar{Z}_{l,t_k}^{N,R}$ ,  $1 \leq l \leq d$  on the function bases  $(p_{l,k})_{0 \leq l \leq d}$ . We denote by  $K_{l,k}$  the size of  $p_{l,k}$ .

# Discretization and simulation of the decoupled FBSDE

## Step 4. Monte-Carlo Simulations.

- We simulate  $M$  independent Monte Carlo simulations of  $(\Lambda_{t_k}^N)_{0 \leq k \leq N}$  and  $(\Delta W_k)_{0 \leq k \leq N-1}$ . We denote  $(\Lambda_{t_k}^{N,m})_{1 \leq m \leq M, 0 \leq k \leq N}$  and  $(\Delta W_k^m)_{1 \leq m \leq M, 0 \leq k \leq N-1}$  these simulations.
- We write  $p_{l,k}(\Lambda_{t_k}^{N,m}) = p_{l,k}^m$  and  $B_{l,k}^M$  the matrix of size  $M \times K_{l,k}$  which rows are  $(p_{l,k}^m)^*$ . we denote by  $K_{l,k}^M$  the rank of  $B_{l,k}^M$ .

**Step 5. Truncations.** We force our approximation of  $\bar{Y}^{N,R}$  and  $\bar{Z}^{N,R}$  to be bounded by  $C_y(R)$  and  $C_z(R)$ . For a function  $\psi$ , we define the truncations by  $[\psi]_y(x) = -C_y(R) \vee \psi(x) \wedge C_y(R)$  and  $[\psi]_z(x) = -C_z(R) \vee \psi(x) \wedge C_z(R)$

# The algorithm

$$\rightarrow \bar{y}_{t_N}^{N,R,M} = 1$$

$\rightarrow$  iteration: for  $k = N - 1, \dots, 0$ , assume that  $\bar{y}^{N,R,M}(\Lambda_{t_{k+1}}^N)$  is known, then we solve the  $d + 1$  least-squares problems:

$$\alpha_{l,k}^M = \arg \inf_{\alpha} \frac{1}{M} \sum_{m=1}^M |\bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}) \frac{[\Delta W_{l,k}^m]_w}{h} - \alpha \cdot p_{l,k}^m|^2$$

$$\begin{aligned} \alpha_{0,k}^M &= \arg \inf_{\alpha} \frac{1}{M} \sum_{m=1}^M |\bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}) \\ &+ h \bar{g}^R(\Lambda_{t_k}^{N,m}, \bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}), [\alpha_{l,k}^M \cdot p_{l,k}^m]_z) - \alpha \cdot p_{0,k}^m|^2 \end{aligned}$$

$$\rightarrow \bar{y}_k^{N,R,M}(\cdot) = [\alpha_{0,k}^M \cdot p_{0,k}]_y(\cdot) \text{ and } \bar{z}_k^{N,R,M}(\cdot) = [\alpha_{l,k}^M \cdot p_{l,k}]_z(\cdot).$$

# The algorithm

**Assumption (H4)**  $(p_{l,k})_{0 \leq l \leq d, 0 \leq k \leq N}$  is a complete orthonormal system with respect to the empirical scalar product  $\langle \cdot, \cdot \rangle_{k,M}$

$$\langle \psi_1, \psi_2 \rangle_{k,M} = \frac{1}{M} \sum_{m=1}^M \psi_1(\Lambda_{t_k}^{N,m}) \psi_2(\Lambda_{t_k}^{N,m})$$

$$\rightarrow \frac{(B_{l,k}^M)^* (B_{l,k}^M)}{M} = Id$$

$$\alpha_{l,k}^M = \frac{1}{M} \sum_{m=1}^M p_{l,k}^m \bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}) \frac{[\Delta W_{l,k}^m]_w}{h}$$

$$\alpha_{0,k}^M = \frac{1}{M} \sum_{m=1}^M p_{0,k}^m \left( \bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}) \right.$$

$$\left. + h \bar{g}^R(\Lambda_{t_k}^{N,m}, \bar{y}_{k+1}^{N,R,M}(\Lambda_{t_{k+1}}^{N,m}), [\alpha_{l,k}^M \cdot p_{l,k}^m] z) \right)$$

# The algorithm

- For the convergence, we refer to Lemor, Gobet Warin (2006)
- The error induced by the localization has a contribution of order  $h$
- We choose the hypercube basis. We choose a domain  $D$  and we partition it into small hypercubes of edge  $\kappa$
- To get a global (squared) error of order  $h^\beta$ ,  $\beta \in (0, 1]$ , we have to choose  $\kappa \approx h^{\frac{\beta+1}{2}}$  (i.e.  $K \approx Ch^{-\frac{d(\beta+1)}{2}}$ ) and  $M \approx Ch^{-d(\beta+1)-(\beta+2)} \log(h^{-\frac{d(\beta+1)}{4}-\frac{\beta+1}{2}})$

# Numerical Approach

We simulate the  $d$ -dimensional stochastic factor  $\Lambda = (\Lambda_t)_{t \in [0, T]}$  by using Monte-Carlo method. We denote by  $\Lambda_{t_k}^{N, m} \in \mathbb{R}^d$  the  $m$ -th approximation of  $\Lambda_{t_k}$ . We determine  $d_{i, k}^{max} = \max_{1 \leq m \leq M} \Lambda_{i, t_k}^{N, m}$  and  $d_{i, k}^{min} = \min_{1 \leq m \leq M} \Lambda_{i, t_k}^{N, m}$ . We consider

$$D_k = \left\{ \lambda \in \mathbb{R}^d, \lambda_i \in [d_{i, k}^{min}, d_{i, k}^{max}], \text{ s.t. for all } 1 \leq i \leq d \right\}.$$

We partition into small hypercubes of edge  $\kappa$  the domain  $D_k$ .

## One risky asset model results

$$\begin{aligned}\frac{dS_t}{S_t} &= bdt + \sigma(\Lambda_t)dW_t + \bar{\sigma}(\Lambda_t)d\bar{W}_t, \\ d\Lambda_t &= (a - \theta\Lambda_t)dt + dW_t, \Lambda_0 = 0,\end{aligned}$$

where  $\sigma(\lambda) = \epsilon_1 + \frac{a_1}{\sqrt{\lambda^2 + \beta_1^2}}$  and  $\bar{\sigma}(\lambda) = \epsilon_2 + \frac{a_2}{\sqrt{\lambda^2 + \beta_2^2}}$ .

Table 1 gives the values of our model parameters.

**Table:** Values for the model's parameters

$d$	$T$	$b$	$\epsilon_1$	$\epsilon_2$	$a_1$
1	1	0.03	$0.71 \times 10^{-6}$	$0.704 \times 10^{-6}$	0.71
$a_2$	$\beta_1$	$\beta_2$	$a$	$\theta$	$\delta$
0.704	4	3	0.05	1	0.5

# One risky asset model results

We determine  $\bar{Y}_0$ , then  $v_1(0, x, \Lambda_0) := \frac{x^\delta}{\delta} \exp(-\bar{Y}_0)$ . We

compute  $v_2(0, x, \Lambda_0) := \frac{1}{M} \sum_{m=1}^M U(X_T^m)$ , where

$$X_T^m = x + \sum_{k=0}^{N-1} X_{t_k}^m \hat{\pi}_{t_k}^m \frac{S_{t_{k+1}}^m - S_{t_k}^m}{S_{t_k}^m}.$$

**Table:** Results for the value function

$M$	4000	6000	8000	10000
$v_1(0, 10, \Lambda_0)$	6.375	6.375	6.375	6.362
$v_2(0, 10, \Lambda_0)$	6.321	6.346	6.367	6.361

## Two risky assets model results

$$\frac{dS_t^1}{S_t^1} = b_1 dt + \nu_1(\Lambda_{1t})(\rho_1 dW_t^1 + \sqrt{1 - \rho_1^2} d\bar{W}_t^1),$$

$$\begin{aligned} \frac{dS_t^2}{S_t^2} &= b_2 dt + \nu_2(\Lambda_{2t}) \left( \rho(\rho_1 dW_t^1 + \sqrt{1 - \rho_1^2} d\bar{W}_t^1) \right. \\ &\quad \left. + \sqrt{1 - \rho^2}(\rho_2 dW_t^2 + \sqrt{1 - \rho_2^2} d\bar{W}_t^2) \right), \end{aligned}$$

$$d\Lambda_{1t} = (a_1 - \theta_1 \Lambda_{1t}) dt + dW_t^1, \quad \Lambda_{10} = 0,$$

$$d\Lambda_{2t} = (a_2 - \theta_2 \Lambda_{2t}) dt + dW_t^2, \quad \Lambda_{20} = 0,$$

where  $\nu_i(\lambda) = \epsilon_i + \frac{1}{\sqrt{\lambda^2 + \beta_i^2}}$ ,  $i \in \{1, 2\}$ . Our model parameters are given in Table 3.

## Two risky assets model results

**Table:** Values for the model's parameters

$d$	$T$	$h$	$b_1$	$b_2$	$a_1$	$\theta_1$	$a_2$	
2	0.4	0.04	0.05	0.03	0.75	1	0.5	
$\theta_2$	$\rho$	$\rho_1$	$\rho_2$	$\delta$	$\epsilon_1$	$\epsilon_2$	$\beta_1$	$\beta_2$
1	0.3	0.75	0.5	0.5	$10^{-6}$	$10^{-6}$	4	3

We determine  $\bar{Y}_0$ , then  $v_1(0, x, \Lambda_0) := \frac{x^\delta}{\delta} \exp(-\bar{Y}_0)$ .

## Two risky assets model results

Table: Results for the value function

$M$	6000	8000	9000
$v_1(0, 10, \Lambda_0)$	6.376	6.372	6.372

The optimal investment strategy at  $t = 0$  is given by  $\hat{\pi}_0^* = (0, 0.91, 0.09)$ . It is optimal to invest only in risky assets. It is better to invest more in the first risky asset since the drift of the second is higher and the level of the volatility of  $S^1$  is lower ( $\beta_1 > \beta_2$ ).