

# Risk-Averse Control of Discrete-Time Partially Observable Systems

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## How to Measure Risk of Sequences?

Probability space  $(\Omega, \mathcal{F}, P)$  with filtration  $\mathcal{F}_1 \subset \dots \subset \mathcal{F}_T \subset \mathcal{F}$

Adapted sequence of random variables (costs)  $Z_1, Z_2, \dots, Z_T$

Spaces:  $\mathcal{Z}_t = \mathcal{L}_{\bar{s}}(\Omega, \mathcal{F}_t, P)$ ,  $\bar{s} \in [1, \infty)$ , and  $\mathcal{Z}_{t,T} = \mathcal{Z}_t \times \dots \times \mathcal{Z}_T$

### Conditional Risk Measure

A mapping  $\rho_{t,T} : \mathcal{Z}_{t,T} \rightarrow \mathcal{Z}_t$  satisfying the **monotonicity condition**:

$$\rho_{t,T}(Z) \leq \rho_{t,T}(W) \text{ for all } Z, W \in \mathcal{Z}_{t,T} \text{ such that } Z \leq W$$

### Dynamic Risk Measure

A sequence of conditional risk measures  $\rho_{t,T} : \mathcal{Z}_{t,T} \rightarrow \mathcal{Z}_t$ ,  $t = 1, \dots, T$

$$\rho_{1,T}(Z_1, Z_2, Z_3, \dots, Z_T) \in \mathcal{Z}_1 = \mathbb{R}$$

$$\rho_{2,T}(Z_2, Z_3, \dots, Z_T) \in \mathcal{Z}_2$$

$$\rho_{3,T}(Z_3, \dots, Z_T) \in \mathcal{Z}_3$$

$\vdots$

# Time Consistency of Dynamic Risk Measures

A dynamic risk measure  $\{\rho_{t,T}\}_{t=1}^T$  is **time-consistent** if for all  $\tau < \theta$

$$Z_k = W_k, \quad k = \tau, \dots, \theta - 1 \quad \text{and} \quad \rho_{\theta,T}(Z_{\theta}, \dots, Z_T) \leq \rho_{\theta,T}(W_{\theta}, \dots, W_T)$$

imply that  $\rho_{\tau,T}(Z_{\tau}, \dots, Z_T) \leq \rho_{\tau,T}(W_{\tau}, \dots, W_T)$

Define  $\rho_t(Z_{t+1}) = \rho_{t,T}(0, Z_{t+1}, 0, \dots, 0)$

## Nested Decomposition Theorem

Suppose a dynamic risk measure  $\{\rho_{t,T}\}_{t=1}^T$  is time-consistent, and

$$\rho_{t,T}(0, \dots, 0) = 0$$

$$\rho_{t,T}(Z_t, Z_{t+1}, \dots, Z_T) = Z_t + \rho_{t,T}(0, Z_{t+1}, \dots, Z_T)$$

Then for all  $t$  we have the representation

$$\rho_{t,T}(Z_t, \dots, Z_T) = Z_t + \rho_t \left( Z_{t+1} + \rho_{t+1} \left( Z_{t+2} + \dots + \rho_{T-1}(Z_T) \right) \dots \right)$$

- State space  $\mathcal{X}$  (Borel)
- Control space  $\mathcal{U}$  (Borel)
- State-control histories
$$g_t = (x_1, u_1, \dots, x_{t-1}, u_{t-1}, x_t) \in \mathcal{G}_t = (\mathcal{X} \times \mathcal{U})^{t-1} \times \mathcal{X}$$
- Feasible control sets  $U_t : \mathcal{G}_t \rightrightarrows \mathcal{U}$ ,  $t = 1, 2, \dots$
- Controlled transition kernels  $Q_t : \text{graph}(U_t) \rightarrow \mathcal{P}(\mathcal{X})$ ,  $t = 1, 2, \dots$   
 $\mathcal{P}(\mathcal{X})$  - set of probability measures on  $\mathcal{X}$
- Cost functions  $c_t : \mathcal{X} \times \mathcal{U} \rightarrow \mathbb{R}$ ,  $t = 1, 2, \dots$
- State history  $h_t = (x_1, \dots, x_t) \in \mathcal{X}^t$  (up to time  $t = 1, 2, \dots$ )
- Policy  $\pi_t : \mathcal{X}^t \rightarrow \mathcal{U}$ ,  $t = 1, 2, \dots$

For any policy  $\pi$ , the transition kernels can be rewritten as measurable functions from  $\mathcal{X}^t$  to  $\mathcal{P}(\mathcal{X})$ :

$$Q_t^\pi : (x_1, \dots, x_t) \mapsto Q_t(x_1, \pi_1(x_1), \dots, x_t, \pi_t(x_1, \dots, x_t))$$

Finite horizon expected cost problem:

$$\min_{\pi_1, \pi_2, \dots, \pi_T} E^{\pi} \left[ \sum_{t=1}^T c_t(x_t, u_t) \right]$$

with controls  $u_t = \pi_t(x_1, \dots, x_t)$

Standard Results for Markov systems:

- A **deterministic Markov policy** is optimal
- Optimal policy can be found by **dynamic programming equations**
- In a **partially observable case**, history dependence can be handled via **belief state** and **Bayes operator** to reduce to Markov model

Our Intention

Introduce **risk aversion** to the problem by replacing the expected value by **dynamic risk measures**

# Using Dynamic Risk Measures for Controlled Processes

- Controlled process  $x_t^\pi$ ,  $t = 1, \dots, T$
- Policy  $\pi = \{\pi_1, \dots, \pi_T\}$  with  $u_t = \pi_t(x_1, \dots, x_t)$  implies measure  $P^\pi$
- Cost sequence  $Z_t^\pi = c(x_t, u_t)$  (bounded),  $t = 1, \dots, T$ ,
- **Dynamic time-consistent risk measure**

$$J_T(\pi) = Z_1^\pi + \rho_1^\pi(Z_2^\pi + \dots + \rho_{T-1}^\pi(Z_T^\pi) \dots)$$

- Risk-averse optimal control problem:  $\min_{\pi} J_T(\pi)$

## Difficulties

- Probability measure  $P^\pi$ , processes  $x_t^\pi$  and  $Z_t^\pi$  depend on policy  $\pi$
- Risk measures  $\rho_t^\pi(\cdot)$  depend on  $\pi$  and history;  
no Markov policies for Markov models

## Idea

We only need to measure risk of random sequences that may occur

History  $h_t = (x_1, \dots, x_t)$ . Process  $Z_t^\pi(h_t) = c(x_t, \pi_t(h_t))$ ,  $t = 1, \dots, T$

A family of conditional risk measures  $\{\rho_{t,T}^\pi\}_{t=1,\dots,T}^{\pi \in \Pi}$  is **stochastically conditionally time-consistent** if for all feasible policies  $\pi, \pi'$ , all  $1 \leq t \leq T-1$ , and for all histories  $h_t \in \mathcal{X}^t$ , the relations

$$Z_t^\pi(h_t) = Z_t^{\pi'}(h_t)$$

$$(\rho_{t+1,T}^\pi(Z_{t+1}^\pi, \dots, Z_T^\pi) | H_t^\pi = h_t) \preceq_{\text{st}} (\rho_{t+1,T}^{\pi'}(Z_{t+1}^{\pi'}, \dots, Z_T^{\pi'}) | H_t^{\pi'} = h_t)$$

imply

$$\rho_{t,T}^\pi(Z_t^\pi, \dots, Z_T^\pi)(h_t) \leq \rho_{t,T}^{\pi'}(Z_t^{\pi'}, \dots, Z_T^{\pi'})(h_t)$$

The conditional stochastic order  $\preceq_{\text{st}}$ :

$$\begin{aligned} Q_t^\pi(h_t) &(\{y : Z_t^\pi(h_t) + \rho_{t+1,T}^\pi(Z_{t+1}^\pi, \dots, Z_T^\pi)(h_t, y) > \eta\}) \\ &\leq Q_t^{\pi'}(h_t) (\{y : Z_t^{\pi'}(h_t) + \rho_{t+1,T}^{\pi'}(Z_{t+1}^{\pi'}, \dots, Z_T^{\pi'})(h_t, y) > \eta\}) \end{aligned}$$

A family of process-based dynamic risk measures  $\{\rho_{t,T}^{\pi}\}_{t=1,\dots,T}^{\pi \in \Pi}$  for a Markov decision problem is **Markovian** if for all Markov policies  $\pi \in \Pi$ , for any measurable and bounded  $c_1, \dots, c_T : \mathcal{X} \times \mathcal{U} \rightarrow \mathbb{R}$ , and for all  $h_t = (x_1, \dots, x_t)$  and  $h'_t = (x'_1, \dots, x'_t)$  such that  $x_t = x'_t$ , we have

$$\begin{aligned} \rho_{t,T}^{\pi}(c_t(X_t, \pi_t(X_t)), \dots, c_T(X_T, \pi_T(X_T)))(h_t) \\ = \rho_{t,T}^{\pi}(c_t(X_t, \pi_t(X_t)), \dots, c_T(X_T, \pi_T(X_T)))(h'_t). \end{aligned}$$

If the current state  $x_t$  is the same, and the same Markov policy  $\pi$  is used, then the risk is the same. The risk measure can be written as a function of the state:

$$\rho_{t,T}^{\pi}(c_t(X_t, \pi_t(X_t)), \dots, c_T(X_T, \pi_T(X_T)))(x_t)$$



For a fixed history-dependent policy  $\Pi$  and every  $h_t \in \mathcal{X}^t$ , we write

$$v_t^\pi(h_t) = \rho_{t,T}^\pi(c_t(X_t, \pi_t(H_t)), \dots, c_T(X_T, \pi_T(H_T)))(h_t)$$

If a family of process-based dynamic risk measures  $\{\rho_{t,T}^\pi\}_{t=1,\dots,T}^{\Pi \in \Pi}$  is translation-invariant and stochastically conditionally time-consistent, then there exist **transition risk mappings**

$$\sigma_t : \left\{ \bigcup_{\pi \in \Pi} \text{graph}(Q_t^\pi) \right\} \times \mathcal{V} \rightarrow \mathbb{R},$$

( $\mathcal{V}$  - space of measurable bounded functions on  $\mathcal{X}$ )

such that for all  $\pi$ , for all  $t$ , and all  $h_t$ , the functional  $\sigma_t(h_t, Q_t^\pi(h_t), \cdot)$  is a **law-invariant risk measure** on  $(\mathcal{X}, \mathcal{B}(\mathcal{X}), Q_t^\pi)$  and we have

$$v_t^\pi(h_t) = c_t(x_t, \pi_t(h_t)) + \sigma_t(h_t, Q_t^\pi(h_t), v_{t+1}^\pi(h_t, \cdot)), \quad t = 1, \dots, T-1$$

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If a family of process-based dynamic risk measures  $\{\rho_{t,T}^\pi\}_{t=1,\dots,T}^{\pi \in \Pi}$  is Markovian, translation-invariant, and stochastically conditionally time-consistent, then there exist **transition risk mappings**

$$\sigma_t : \{(x, Q_t(x, u)) : u \in U(x), x \in \mathcal{X}\} \times \mathcal{V} \rightarrow \mathbb{R}, \quad t = 1, \dots, T-1$$

( $\mathcal{V}$  - space of measurable bounded functions on  $\mathcal{X}$ )

such that for all  $\pi \in \Pi$ , for all  $t = 1, \dots, T-1$ , and all  $h_t \in \mathcal{X}^t$ , the functional  $\sigma_t(x_t, Q_t(x_t, \pi_t(h_t), \cdot))$  is a **law-invariant risk measure** on  $(\mathcal{X}, \mathcal{B}(\mathcal{X}), Q_t)$  and

$$v_t^\pi(h_t) = c_t(x_t, \pi_t(h_t)) + \sigma_t(x_t, Q_t(x_t, \pi_t(h_t)), v_{t+1}^\pi(h_t, \cdot)), \quad t = 1, \dots, T-1$$

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$$v_t^\pi(x_t) = c_t(x_t, \pi_t(x_t)) + \sigma_t(x_t, Q_t(x_t, \pi_t(x_t)), v_{t+1}^\pi(\cdot)), \quad t = 1, \dots, T-1$$

## Finite Horizon Risk-Averse Control Problem

Consider a controlled **Markov** process  $\{X_t\}$  with  $u_t = \pi_t(X_1, \dots, X_t)$ .

Risk-averse optimal control problem:

$$\min_{\pi} J_T(\pi) = c_1(x_1, u_1) + \rho_1^{\pi} \left( c_2(X_2, u_2) + \dots + \rho_{T-1}^{\pi} \left( c_T(X_T, u_T) \right) \dots \right)$$

### Theorem

If the conditional measures  $\rho_t^{\pi}$  are Markovian (+ general conditions), then the optimal solution is given by the **dynamic programming equations**:

$$v_T(x) = \min_{u \in U_T(x)} c_T(x, u), \quad x \in \mathcal{X}$$

$$v_t(x) = \min_{u \in U_t(x)} \left\{ c_t(x, u) + \sigma_t(x, Q_t(x, u), v_{t+1}) \right\}, \quad x \in \mathcal{X}, \quad t = T-1, \dots, 1.$$

Optimal **Markov policy**  $\hat{\Pi} = \{\hat{\pi}_1, \dots, \hat{\pi}_T\}$  - the minimizers above

- Markov Process:  $\{X_t, Y_t\}_{t=1, \dots, T}$  on the state space  $\mathcal{X} \times \mathcal{Y}$
- The process  $\{Y_t\}$  is observable, while  $\{X_t\}$  is not observable
- Control sets:  $U_t : \mathcal{Y} \rightrightarrows \mathcal{U}$ ,  $t = 1, \dots, T$ .
- Transition kernel:

$$\mathbb{P}[(X_{t+1}, Y_{t+1}) \in C \mid x_1, y_1, u_1, \dots, x_t, y_t, u_t] = Q_t(x_t, y_t, u_t)(C),$$

- **Belief State:** Conditional distribution of  $X_t$  given initial distribution  $\xi_1$  and history  $g_t = (\xi_1, y_1, u_1, y_2, \dots, u_{t-1}, y_t)$

$$[\mathcal{E}_t(g_t)](A) = \mathbb{P}[X_t \in A \mid g_t], \quad \forall A \in \mathcal{B}(\mathcal{X}), \quad t = 1, \dots, T$$

Conditional distribution of the observable part:

$$\mathbb{P}[Y_{t+1} \in B \mid g_t, u_t] = \int_{\mathcal{X}} [Q_t^Y(\cdot, y_t, u_t)](B) d\mathcal{E}_t(g_t),$$

where  $Q_t^Y(x_t, y_t, u_t)$  is the marginal of  $Q_t(x_t, y_t, u_t)$  on the space  $\mathcal{Y}$

Transition of the belief state:

$$\mathcal{E}_{t+1}(g_{t+1}) = \Phi_t(y_t, \mathcal{E}_t(g_t), u_t, y_{t+1})$$

Bayes operator with kernels given by density functions

If  $\mathcal{X} = \{x^1, \dots, x^n\}$  is finite and kernels  $Q_t(x, y, u)$  have density  $q_t(x', \cdot | x, y, u)$  with respect to a finite measure  $\mu_Y$  on  $\mathcal{Y}$ , then

$$[\Phi_t(y, \xi, u, y')](\{x^k\}) = \frac{\sum_{i=1}^n q_t(x^k, y' | x^i, y, u) \xi^i}{\sum_{\ell=1}^n \sum_{i=1}^n q_t(x^\ell, y' | x^i, y, u) \xi^i}$$

Further specialization for **hidden Markov models** is possible, where

- The unobserved process  $\{X_t\}_{t=1, \dots, T}$  is a Markov process itself
- For all  $t = 1, \dots, T$ , the observable part  $Y_t$  is a noisy function of  $X_t$

Extended state history (including belief states):

$$h_t = (y_1, \xi_1, y_2, \xi_2, \dots, y_t, \xi_t) \in \mathbb{H}_t$$

Policies  $\pi = (\pi_1, \dots, \pi_T)$  with decision rules  $\pi_t(h_t) \in U_t(y_t)$

### Markov Policy

For all  $h_t, h'_t \in \mathbb{H}_t$ , if  $y_t = y'_t$  and  $\xi_t = \xi'_t$ , then  $\pi_t(h_t) = \pi_t(h'_t) = \pi_t(y_t, \xi_t)$

Denote  $v_t^\pi(h_t) = \rho_{t,T}^\pi(c_t(Y_t, \pi_t(H_t)), \dots, c_T(Y_T, \pi_T(H_T)))(h_t)$ .

A family of dynamic risk measures  $\{\rho_{t,T}^\pi\}_{t=1, \dots, T}^{\pi \in \Pi}$  is **Markovian** if for any Markov policy  $\pi \in \Pi$ , all  $t = 1, \dots, T$ , and all  $h_t, h'_t \in \mathbb{H}_t$  such that  $y_t = y'_t$  and  $\xi_t = \xi'_t$ , we have

$$v_t^\pi(h_t) = v_t^\pi(h'_t)$$

Thus for a Markovian policy  $\Pi$ , we can write  $v_t^\pi(y_t, \xi_t)$  instead of  $v_t^\pi(h_t)$

We denote the conditional distributions of  $Y_{t+1}$ , given the observations up to time  $t$ , by  $\Upsilon_t(g_t, u_t)$ .

If a family of dynamic risk measures  $\{\rho_{t,T}^\pi\}_{t=1,\dots,T}^{\pi \in \Pi}$  is translation-invariant, stochastically conditionally time-consistent and Markov, then functionals  $\sigma_t : \{(y, \xi, \Upsilon_t(y, \xi, u)) \mid y \in \mathcal{Y}, \xi \in \mathcal{P}(\mathcal{X}), u \in \mathcal{U}_t(y)\} \times \mathcal{P}(\mathcal{Y}) \rightarrow \mathbb{R}$  exist such that for all  $t = 1, \dots, T-1$ , for all  $(y_t, \xi_t) \in \mathcal{Y} \times \mathcal{P}(\mathcal{X})$ , the functional  $\sigma_t(y_t, \xi_t, \cdot, \cdot)$  satisfies the law invariance property, and for all  $\pi$  all  $h_t \in \mathbb{H}_t$ , we have

$$v_t^\pi(h_t) = c_t(y_t, \pi_t(h_t)) + \sigma_t\left(y_t, \xi_t, \Upsilon_t(y_t, \xi_t, \pi_t(h_t)), y' \mapsto v_{t+1}^\pi(h_t, y', \Phi_t(y_t, \xi_t, \pi_t(h_t), y'))\right)$$

The risk model can be formulated in terms of static risk measures on the space of functions of  $y$ , rather than functions of  $(y, \xi)$ .



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$$v_t^\pi(y_t, \xi_t) = c_t(y_t, \pi_t(y_t, \xi_t)) + \sigma_t\left(y_t, \xi_t, \Upsilon_t(y_t, \xi_t, \pi_t(y_t, \xi_t)), y' \mapsto v_{t+1}^\pi(y', \Phi_t(y_t, \xi_t, \pi_t(y_t, \xi_t), y'))\right)$$

The risk model can be formulated in terms of static risk measures on the space of functions of  $y$ , rather than functions of  $(y, \xi)$ .

Risk-averse optimal control problem:

$$\min_{\pi} \rho_{1,T}^{\pi} \{c_1(y_1, u_1), c_2(y_2, u_2), \dots, c_T(y_T, u_T)\}$$

## Theorem

If the risk measure is Markovian (+ general conditions), then the optimal solution is given by the **dynamic programming equations**:

$$v_T^*(y, \xi) = \min_{u \in \mathcal{U}_T(x)} c_T(y, u), \quad y \in \mathcal{Y}, \quad \xi \in \mathcal{P}(\mathcal{X})$$

$$v_t^*(y, \xi) = \min_{u \in \mathcal{U}_t(y)} \left\{ c_t(y, u) + \sigma_t \left( y, \xi, \int_{\mathcal{Y}} K_t^X(x, \cdot, u) d\xi, y' \mapsto v_{t+1}^*(y', \Phi_t(x, \xi, u, y')) \right) \right\},$$
$$y \in \mathcal{Y}, \quad \xi \in \mathcal{P}(\mathcal{X}), \quad t = T-1, \dots, 1$$

Optimal **Markov policy**  $\hat{\Pi} = \{\hat{\pi}_1, \dots, \hat{\pi}_T\}$  - the minimizers above

- In stages  $t = 1, \dots, T$  of the trials successive patients are given drugs (cytotoxic agents), to which **severe toxic response (death)** is possible
- Probability of toxic response depends on the unknown **optimal dose  $\eta^*$**  and the **administered dose (control)  $u$** :

$$F(u, \eta) = \frac{1}{1 - e^{-\varphi(u, \eta)}}$$

- The “**belief state**”  $s_t$ , the probability distribution of the unknown optimal dose, is the current **state of the system**
- The state evolves according to **Bayes operator**, depending on the response of the patient: for  $\eta \in \mathcal{X}$  (the range of doses)

$$s_{t+1}(\eta) \sim \begin{cases} F(u_t, \eta) s_t(\eta) & \text{if toxic} \\ (1 - F(u_t, \eta)) s_t(\eta) & \text{if not toxic} \end{cases}$$

- **Cost per stage**:  $c_t(s_t, u_t) = \gamma_t \mathbb{E}_{s_t}[|u_t - \eta|]$  (other forms possible)

**Medical ethics** naturally motivates **risk-averse control**

# Total Cost Models

Find the best policy  $\Pi = (\pi_1, \dots, \pi_T)$  to determine doses  $u_t = \pi_t(s_t)$

## Expected Value Model

$$\min_{\Pi} \mathbb{E} \left[ \sum_{t=1}^{T+1} \gamma_t |u_t - \eta^*| \right]$$

$\gamma_{T+1}$  is the weight of the **final recommendation**  $u_{T+1}$

## Risk-Averse Value Model

$$\min_{\Pi} \rho_{1, T+1} \left[ \sum_{t=1}^{T+1} \gamma_t |u_t - \eta^*| \right]$$

## Two sources of risk

- Unknown state  $\eta^*$  (only distribution  $s_t$  available)
- Unknown evolution due to random response  $r_t$

At each time  $t$ , assume that this is the last test before the final recommendation, and solve the two-stage problem

### Risk-Neutral

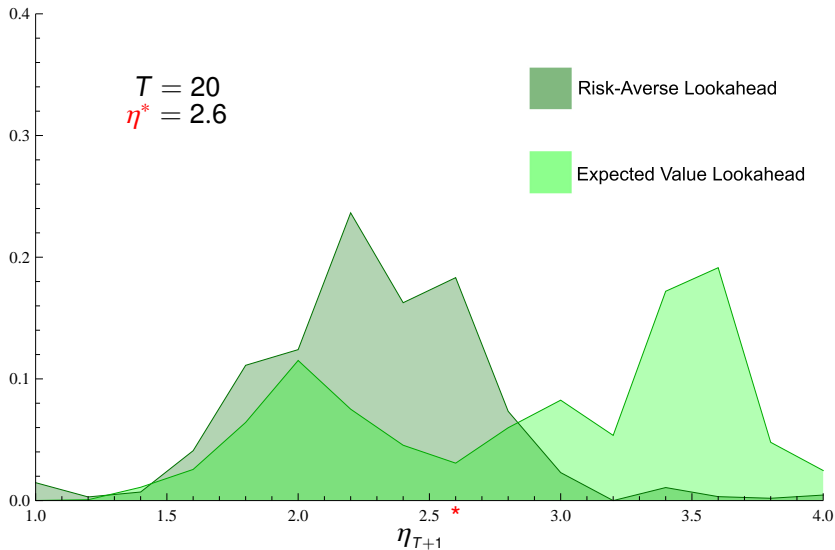
$$\min_{u_t} \mathbb{E}_{s_t} \left\{ \gamma_t |u_t - \eta| + \bar{\gamma}_{t+1} \mathbb{E}_{r_t} \left[ \min_{u_{t+1}} \mathbb{E}_{s_{t+1}} |u_{t+1} - \eta| \right] \right\}$$

### Risk-Averse

$$\min_{u_t} \rho_{s_t} \left\{ \gamma_t |u_t - \eta| + \bar{\gamma}_{t+1} \rho_{r_t} \left[ \min_{u_{t+1}} \mathbb{E}_{s_{t+1}} |u_{t+1} - \eta| \right] \right\}$$

$$\bar{\gamma}_{t+1} = \sum_{\tau=t+1}^{T+1} \gamma_{\tau} \quad (\text{weight of the future})$$

## Distribution of Dosage



- A. Ruszczyński, Risk-averse dynamic programming for Markov decision processes, *Mathematical Programming, Series B* 125 (2010) 235–261
- Ö. Çavuş and A. Ruszczyński, Computational methods for risk-averse undiscounted transient Markov models, *Operations Research*, 62 (2), 2014, 401–417.
- Ö. Çavuş and A. Ruszczyński, Risk-averse control of undiscounted transient Markov models, *SIAM J. on Control and Optimization*, 52(6), 2014, 3935–3966.
- J. Fan and A. Ruszczyński, Process-based risk measures for observable and partially observable discrete-time controlled systems, *submitted for publication*.
- C. McGinity, D. Dentcheva and A. Ruszczyński, Risk-averse approximate dynamic programming for partially observable systems with application to clinical trials, *in preparation*.