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A mathematical introduction to molecular dynamics

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## Outline

- Sampling high-dimensional probability measures
  - Statistical physics
  - Bayesian inference
  - Standard sampling techniques (low-dimensional)

- Sampling with stochastic differential equations
  - A primer on SDEs
  - Langevin-like dynamics

• Variance reduction

## General references (1)

- Computational Statistical Physics
  - D. Frenkel and B. Smit, Understanding Molecular Simulation, From Algorithms to Applications (Academic Press, 2002)
  - M. Tuckerman, *Statistical Mechanics: Theory and Molecular Simulation* (Oxford, 2010)
  - M. P. Allen and D. J. Tildesley, *Computer simulation of liquids* (Oxford University Press, 1987)
  - D. C. Rapaport, *The Art of Molecular Dynamics Simulations* (Cambridge University Press, 1995)
  - T. Schlick, Molecular Modeling and Simulation (Springer, 2002)
- Computational Statistics [my personal references... many more out there!]
  - J. Liu, Monte Carlo strategies in scientific computing, Springer, 2008
  - W. R. Gilks, S. Richardson and D. J. Spiegelhalter (eds), *Markov chain Monte Carlo in practice* (Chapman & Hall, 1996)
- Machine learning and sampling
  - C. Bishop, Pattern Recognition and Machine Learning (Springer, 2006)

## General references (2)

- Sampling the canonical measure
  - L. Rey-Bellet, Ergodic properties of Markov processes, *Lecture Notes in Mathematics*, **1881** 1–39 (2006)
  - E. Cancès, F. Legoll and G. Stoltz, Theoretical and numerical comparison of some sampling methods, *Math. Model. Numer. Anal.* 41(2) (2007) 351-390
  - T. Lelièvre, M. Rousset and G. Stoltz, *Free Energy Computations: A Mathematical Perspective* (Imperial College Press, 2010)
  - B. Leimkuhler and C. Matthews, *Molecular Dynamics: With Deterministic and Stochastic Numerical Methods* (Springer, 2015).
  - T. Lelièvre and G. Stoltz, Partial differential equations and stochastic methods in molecular dynamics, *Acta Numerica* **25**, 681-880 (2016)
- Convergence of Markov chains
  - S. Meyn and R. Tweedie, *Markov Chains and Stochastic Stability* (Cambridge University Press, 2009)
  - R. Douc, E. Moulines, P. Priouret and P. Soulier, Markov chains (Springer, 2018)

## An introduction to statistical physics

## Statistical physics (1)

- Aims of computational statistical physics
  - numerical microscope
  - computation of average properties, static or dynamic
- Orders of magnitude
  - distances  $\sim 1~{\mathring{A}} = 10^{-10}~{\rm m}$
  - $\bullet$  energy per particle  $\sim k_{\rm B}T \sim 4 \times 10^{-21}~{\rm J}$  at room temperature
  - $\bullet$  atomic masses  $\sim 10^{-26}~{\rm kg}$
  - time  $\sim 10^{-15}~{\rm s}$
  - number of particles  $\sim \mathcal{N}_A = 6.02 imes 10^{23}$

#### • "Standard" simulations

- $10^6$  particles ["world records": around  $10^9$  particles]
- $\bullet$  integration time: (fraction of) ns ["world records": (fraction of)  $\mu s]$

## Statistical physics (2)

#### What is the melting temperature of argon?



## Statistical physics (3)

"Given the structure and the laws of interaction of the particles, what are the macroscopic properties of the matter composed of these particles?"



Equation of state (pressure/density diagram) for argon at T = 300 K

## Statistical physics (4)

What is the structure of the protein? What are its typical conformations, and what are the transition pathways from one conformation to another?



## Statistical physics (5)

• Microstate of a classical system of  ${\cal N}$  particles:

$$(q,p) = (q_1,\ldots,q_N, p_1,\ldots,p_N) \in \mathcal{E}$$

Positions q (configuration), momenta p (to be thought of as  $M\dot{q}$ )

• In the simplest cases,  $\mathcal{E} = \mathcal{D} imes \mathbb{R}^{3N}$  with  $\mathcal{D} = \mathbb{R}^{3N}$  or  $\mathbb{T}^{3N}$ 

• More complicated situations can be considered: molecular constraints defining submanifolds of the phase space

• Hamiltonian  $H(q,p) = E_{kin}(p) + V(q)$ , where the kinetic energy is

$$E_{\rm kin}(p) = \frac{1}{2} p^T M^{-1} p, \qquad M = \begin{pmatrix} m_1 \, {\rm Id}_3 & 0 \\ & \ddots & \\ 0 & & m_N \, {\rm Id}_3 \end{pmatrix}$$

## Statistical physics (6)

- $\bullet$  All the physics is contained in V
  - ideally derived from quantum mechanical computations
  - in practice, empirical potentials for large scale calculations
- An example: Lennard-Jones pair interactions to describe noble gases

$$V(q_1, \dots, q_N) = \sum_{1 \leqslant i < j \leqslant N} v(|q_j - q_i|)$$

$$v(r) = 4\varepsilon \left[ \left(\frac{\sigma}{r}\right)^{12} - \left(\frac{\sigma}{r}\right)^6 \right]$$

$$V(r$$

## Statistical physics (7)

• Macrostate of the system described by a probability measure

Equilibrium thermodynamic properties (pressure,...)

$$\langle \varphi \rangle_{\mu} = \mathbb{E}_{\mu}(\varphi) = \int_{\mathcal{E}} \varphi(q, p) \, \mu(dq \, dp)$$

- Choice of thermodynamic ensemble
  - least biased measure compatible with the observed macroscopic data
  - Volume, energy, number of particles, ... fixed exactly or in average
  - Equivalence of ensembles (as  $N \to +\infty$ )
- Canonical ensemble = measure on (q, p), average energy fixed H

$$\mu_{\rm NVT}(dq\,dp) = Z_{\rm NVT}^{-1}\,{\rm e}^{-\beta H(q,p)}\,dq\,dp$$

with  $\beta = \frac{1}{k_{\rm B}T}$  the Lagrange multiplier of the constraint  $\int_{\mathcal{E}} H \rho \, dq \, dp = E_0$ Gabriel Stotz (ENPC/INRIA) Edinburgh, Feb. 2020 12/47 Another motivation for sampling high-dimensional probability measures

## Bayesian inference (1)

- Data set  $\{y_i\}_{i=1,...,N_{data}}$
- Elementary likelihood P(y|q), with q parameters of probability measure
- A priori distribution of the parameters  $p_{\rm prior}$  (usually not so informative)

#### Aim

Find the values of the parameters  $\boldsymbol{q}$  describing correctly the data: sample

$$\nu(q) \propto p_{\text{prior}}(q) \prod_{i=1}^{N_{\text{data}}} P(y_i|q)$$

• Example of Gaussian mixture model

## Bayesian inference (2)

 $\bullet$  Elementary likelihood approximated by mixture of K Gaussians

$$P(y \mid \theta) = \sum_{k=1}^{K} a_k \sqrt{\frac{\lambda_k}{2\pi}} \exp\left(-\frac{\lambda_k}{2}(y - \mu_k)^2\right)$$

• Parameters  $\theta = (a_1, \dots, a_{K-1}, \mu_1, \dots, \mu_K, \lambda_1, \dots, \lambda_K)$  with

 $\mu_k \in \mathbb{R}, \quad \lambda_k \ge 0, \quad 0 \le a_k \le 1, \quad a_1 + \dots + a_K = 1$ 

- Prior distribution: Random beta model: additional variable
  - uniform distribution of the weights  $a_k$
  - $\mu_k \sim \mathcal{N}\left(M, R^2/4\right)$  with M = mean of data,  $R = \max y_i \min y_i$
  - $\lambda_k \sim \Gamma(\alpha, \beta)$  with  $\beta \sim \Gamma(g, h)$ , g = 0.2 and  $h = 100g/\alpha R^2$

#### Aim

Find the values of the parameters (namely  $\theta$ , and possibly K as well) describing correctly the data

[RG97] S. Richardson and P. J. Green. *J. Roy. Stat. Soc. B*, 1997. [JHS05] A. Jasra, C. Holmes and D. Stephens, Statist. Science, 2005

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## Bayesian inference (3)



Left: Lengths of snappers ( $N_{data} = 256$ ), and a possible fit for K = 3 using the last configuration from the trajectory plotted in the right picture.

**Right:** Typical sampling trajectory, Metropolis/Gaussian random walk with  $(\sigma_q, \sigma_\mu, \sigma_v, \sigma_\beta) = (0.0005, 0.025, 0.05, 0.005).$ 

[IS88] A. J. Izenman and C. J. Sommer, J. Am. Stat. Assoc., 1988.
 [BMY97] K. Basford et al., J. Appl. Stat., 1997

## Bayesian inference (4)



Left: Thickness of Mexican stamps ("Hidalgo stamp data",  $N_{\text{data}} = 485$ ), and two possible fits for K = 3 ("genuine multimodality", solid line: dominant mode).

Right: Typical sampling trajectory

[TSM86] D. Titterington *et al.*, *Statistical Analysis of Finite Mixture Distributions*, 1986. [FS06] S. Frühwirth-Schnatter, *Finite Mixture and Markov Switching Models*, 2006.

## Bayesian inference (5)



Scatter plot of the marginal distribution of  $(\mu_1, \log \lambda_1)$  for the Fish data, for various values of K = 4, 5, 6

# Standard techniques for sampling probability measures

## Standard techniques to sample probability measures (1)

- $\bullet$  The basis is the generation of numbers uniformly distributed in  $\left[0,1\right]$
- Deterministic sequences which look like they are random...
  - Early methods: linear congruential generators ("chaotic" sequences)

$$x_{n+1} = ax_n + b \mod c, \qquad u_n = \frac{x_n}{c-1}$$

- Known defects: short periods, point alignments, etc, which can be (partially) patched by cleverly combining several generators
- More recent algorithms: shift registers, such as Mersenne-Twister  $\rightarrow$  defaut choice in *e.g.* Scilab, available in the GNU Scientific Library
- Randomness tests: various flavors

### Standard techniques to sample probability measures (2)

- Classical distributions are obtained from the uniform distribution by...
  - inversion of the cumulative function  $F(x) = \int_{-\infty}^{x} f(y) \, dy$  (which is an increasing function from  $\mathbb{R}$  to [0, 1])

$$X = F^{-1}(U) \sim f(x) \, dx$$

 $\begin{array}{l} \operatorname{Proof:} \ \mathbb{P}\{a < X \leqslant b\} = \mathbb{P}\{a < F^{-1}(X) \leqslant b\} = \mathbb{P}\{F(a) < U \leqslant F(b)\} = F(b) - F(a) = \int_a^b f(x) \, dx \\ \operatorname{Example:} \ \operatorname{exponential} \ \operatorname{law} \ \operatorname{of} \ \operatorname{density} \ \lambda \mathrm{e}^{-\lambda x} \mathbf{1}_{\{x \geqslant 0\}}, \ F(x) = \mathbf{1}_{\{x \geqslant 0\}} (1 - \mathrm{e}^{-\lambda x}), \ \operatorname{so} \ \operatorname{that} \ X = -\frac{1}{\lambda} \ln U \\ \end{array}$ 

• change of variables: standard Gaussian  $G = \sqrt{-2\ln U_1}\cos(2\pi U_2)$ Proof:  $\mathbb{E}(f(X,Y)) = \frac{1}{2\pi} \int_{\mathbb{R}^2} f(x,y) e^{-(x^2+y^2)/2} dx dy = \int_0^{+\infty} f\left(\sqrt{r}\cos\theta, \sqrt{r}\sin\theta\right) \frac{1}{2} e^{-r/2} dr \frac{d\theta}{2\pi}$ 

#### using the rejection method

Find a probability density g and a constant  $c \ge 1$  such that  $0 \le f(x) \le cg(x)$ . Generate i.i.d. variables  $(X^n, U^n) \sim g(x) \, dx \otimes \mathcal{U}[0, 1]$ , compute  $r^n = \frac{f(X^n)}{cg(X^n)}$ , and accept  $X^n$  if  $r^n \ge U^n$ 

## Standard techniques to sample probability measures (3)

- The previous methods work only
  - for low-dimensional probability measures
  - when the normalization constants of the probability density are known
- In more complex cases, one needs to resort to trajectory averages



- Find methods for which
  - the convergence is guaranteed? (and in which sense?)
  - error estimates are available? (typically with Central Limit Theorem)

## Standard techniques to sample probability measures (4)

• Assume that  $x^n \sim \pi$  are idependently and identically distributed (i.i.d.)

Law of Large Numbers for  $\varphi \in L^1(\pi)$ 

$$S_{N_{\text{iter}}} = \frac{1}{N_{\text{iter}}} \sum_{n=1}^{N_{\text{iter}}} \varphi(x^n) \xrightarrow[N_{\text{iter}} \to +\infty]{} \mathbb{E}_{\pi}(\varphi) = \int_{\mathcal{X}} \varphi \, d\pi \quad \text{almost surely}$$

Central Limit Theorem for  $\varphi \in L^2(\pi)$ 

$$\sqrt{N_{\text{iter}}} \left( S_{N_{\text{iter}}} - \int \varphi \, d\pi \right) \xrightarrow[N_{\text{iter}} \to +\infty]{} \mathcal{N}(0, \sigma_{\varphi}^2), \ \sigma_{\varphi}^2 = \int_{\mathcal{X}} \left[ \varphi - \mathbb{E}_{\pi}(\varphi) \right]^2 \, d\pi$$

• This should be thought of in practice as  $S_{N_{\mathrm{iter}}} \simeq \mathbb{E}_{\pi}(\varphi) + \frac{\sigma_{\varphi}}{\sqrt{N_{\mathrm{iter}}}}\mathcal{G}$ 

# Sampling with stochastic differential equations

## Langevin dynamics

• Stochastic perturbation of the Hamiltonian dynamics : friction  $\gamma > 0$ 

$$\begin{cases} dq_t = M^{-1} p_t \, dt \\ dp_t = -\nabla V(q_t) \, dt - \gamma M^{-1} p_t \, dt + \sqrt{\frac{2\gamma}{\beta}} \, dW_t \end{cases}$$

#### Motivations

- Ergodicity can be proved and is indeed observed in practice
- Many useful extensions

#### • Aims

- Understand the meaning of this equation
- Understand why it samples the canonical ensemble
- Implement appropriate discretization schemes
- Estimate the errors (systematic biases vs. statistical uncertainty)

## An intuitive view of the Brownian motion (1)

• Independant Gaussian increments whose variance is proportional to time

 $\forall 0 < t_0 \leqslant t_1 \leqslant \cdots \leqslant t_n, \qquad W_{t_{i+1}} - W_{t_i} \sim \mathcal{N}(0, t_{i+1} - t_i)$ 

where the increments  $W_{t_{i+1}} - W_{t_i}$  are independent

+  $G\sim \mathcal{N}(m,\sigma^2)$  distributed according to the probability density

$$g(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-m)^2}{2\sigma^2}\right)$$

• The solution of  $dq_t = \sigma dW_t$  can be thought of as the limit  $\Delta t \to 0$ 

$$q^{n+1} = q^n + \sigma \sqrt{\Delta t} G^n, \qquad G^n \sim \mathcal{N}(0, 1) \text{ i.i.d.}$$

where  $q^n$  is an approximation of  $q_{n\Delta t}$ 

- Note that  $q^n \sim \mathcal{N}(q^0, \sigma^2 n \Delta t)$
- Multidimensional case:  $W_t = (W_{1,t}, \dots, W_{d,t})$  where  $W_i$  are independent Gabriel Stoltz (ENPC/INRIA) Edinburgh, Feb. 2020 26/47

## An intuitive view of the Brownian motion (2)

- Analytical study of the process: law  $\psi(t,q)$  of the process at time  $t \rightarrow$  distribution of all possible realizations of  $q_t$  for
  - a given initial distribution  $\psi(0,q)$ , e.g.  $\delta_{q^0}$
  - and all realizations of the Brownian motion

Averages at time t

$$\mathbb{E}\Big(A(q_t)\Big) = \int_{\mathcal{D}} A(q) \,\psi(t,q) \,dq$$

• Partial differential equation governing the evolution of the law

Fokker-Planck equation

$$\partial_t \psi = \frac{\sigma^2}{2} \Delta \psi$$

Here, simple heat equation  $\rightarrow$  "diffusive behavior"

## An intuitive view of the Brownian motion (3)

• Proof: Taylor expansion, beware random terms of order  $\sqrt{\Delta t}$ 

$$A\left(q^{n+1}\right) = A\left(q^{n} + \sigma\sqrt{\Delta t} G^{n}\right)$$
$$= A\left(q^{n}\right) + \sigma\sqrt{\Delta t}G^{n} \cdot \nabla A\left(q^{n}\right) + \frac{\sigma^{2}\Delta t}{2}\left(G^{n}\right)^{T}\left(\nabla^{2}A\left(q^{n}\right)\right)G^{n} + O\left(\Delta t^{3/2}\right)$$

Taking expectations (Gaussian increments  $G^n$  independent from the current position  $q^n$ )

$$\mathbb{E}\left[A\left(q^{n+1}\right)\right] = \mathbb{E}\left[A\left(q^{n}\right) + \frac{\sigma^{2}\Delta t}{2}\Delta A\left(q^{n}\right)\right] + O\left(\Delta t^{3/2}\right)$$
  
Therefore,  $\mathbb{E}\left[\frac{A\left(q^{n+1}\right) - A\left(q^{n}\right)}{\Delta t} - \frac{\sigma^{2}}{2}\Delta A\left(q^{n}\right)\right] \to 0$ . On the other hand,  
 $\mathbb{E}\left[\frac{A\left(q^{n+1}\right) - A\left(q^{n}\right)}{\Delta t}\right] \to \partial_{t}\left(\mathbb{E}\left[A(q_{t})\right]\right) = \int_{\mathcal{D}} A(q)\partial_{t}\psi(t,q)\,dq.$ 

This leads to

$$0 = \int_{\mathcal{D}} A(q) \partial_t \psi(t,q) \, dq - \frac{\sigma^2}{2} \int_{\mathcal{D}} \Delta A(q) \, \psi(t,q) \, dq = \int_{\mathcal{D}} A(q) \left( \partial_t \psi(t,q) - \frac{\sigma^2}{2} \Delta \psi(t,q) \right) dq$$

This equality holds for all observables A.

## General SDEs (1)

 $\bullet$  State of the system  $X\in\mathbb{R}^d$  , m-dimensional Brownian motion, diffusion matrix  $\sigma\in\mathbb{R}^{d\times m}$ 

$$dX_t = b(X_t) dt + \sigma(X_t) dW_t$$

to be thought of as the limit as  $\Delta t \to 0$  of  $(X^n \text{ approximation of } X_{n\Delta t})$ 

$$X^{n+1} = X^n + \Delta t \, b \, (X^n) + \sqrt{\Delta t} \, \sigma(X^n) G^n, \qquad G^n \sim \mathcal{N} \left( 0, \mathrm{Id}_m \right)$$

• Generator

$$\mathcal{L} = b(x) \cdot \nabla + \frac{1}{2}\sigma\sigma^{T}(x) : \nabla^{2} = \sum_{i=1}^{d} b_{i}(x)\partial_{x_{i}} + \frac{1}{2}\sum_{i,j=1}^{d} \left[\sigma\sigma^{T}(x)\right]_{i,j}\partial_{x_{i}}\partial_{x_{j}}$$

• Proceeding as before, it can be shown that

$$\partial_t \Big( \mathbb{E} \left[ A(X_t) \right] \Big) = \int_{\mathcal{X}} A \, \partial_t \psi = \mathbb{E} \Big[ \left( \mathcal{L}A \right) \left( X_t \right) \Big] = \int_{\mathcal{X}} \left( \mathcal{L}A \right) \psi$$

## General SDEs (2)

#### Fokker-Planck equation

$$\partial_t \psi = \mathcal{L}^* \psi$$

where  $\mathcal{L}^*$  is the adjoint of  $\mathcal L$ 

$$\int_{\mathcal{X}} (\mathcal{L}A) (x) B(x) dx = \int_{\mathcal{X}} A(x) (\mathcal{L}^*B) (x) dx$$

Invariant measures are stationary solutions of the Fokker-Planck equation

Invariant probability measure  $\psi_{\infty}(x) dx$ 

$$\mathcal{L}^*\psi_{\infty} = 0, \qquad \int_{\mathcal{X}} \psi_{\infty}(x) \, dx = 1, \qquad \psi_{\infty} \ge 0$$

• When  $\mathcal{L}$  is elliptic (*i.e.*  $\sigma\sigma^T$  has full rank: the noise is sufficiently rich), the process can be shown to be irreducible = accessibility property

$$P_t(x,\mathcal{S}) = \mathbb{P}(X_t \in \mathcal{S} \mid X_0 = x) > 0$$

## General SDEs (3)

- Sufficient conditions for ergodicity
  - irreducibility
  - existence of an invariant probability measure  $\psi_\infty(x)\,dx$

Then the invariant measure is unique and

$$\lim_{T \to \infty} \frac{1}{T} \int_0^T \varphi(X_t) \, dt = \int_{\mathcal{X}} \varphi(x) \, \psi_\infty(x) \, dx \qquad \text{a.s.}$$

• Rate of convergence given by Central Limit Theorem:  $\widetilde{\varphi} = \varphi - \int \varphi \, \psi_{\infty}$ 

$$\sqrt{T} \left( \frac{1}{T} \int_0^T \varphi(X_t) \, dt - \int \varphi \, \psi_\infty \right) \xrightarrow[T \to +\infty]{\text{law}} \mathcal{N}(0, \sigma_\varphi^2)$$

with  $\sigma_{\varphi}^2 = 2 \mathbb{E} \left[ \int_0^{+\infty} \widetilde{\varphi}(X_t) \widetilde{\varphi}(X_0) dt \right]$  (proof: later, discrete time setting)

## SDEs: numerics (1)

- Numerical discretization: various schemes (Markov chains in all cases)
- Example: Euler-Maruyama

$$X^{n+1} = X^n + \Delta t \, b(X^n) + \sqrt{\Delta t} \, \sigma(X^n) \, G^n, \qquad G^n \sim \mathcal{N}(0, \mathrm{Id}_d)$$

• Standard notions of error: fixed integration time  $T<+\infty$ 

• Strong error 
$$\sup_{0 \le n \le T/\Delta t} \mathbb{E} |X^n - X_{n\Delta t}| \le C\Delta t^p$$

- Weak error:  $\sup_{0 \leqslant n \leqslant T/\Delta t} \left| \mathbb{E} \left[ \varphi \left( X^n \right) \right] \mathbb{E} \left[ \varphi \left( X_{n\Delta t} \right) \right] \right| \leqslant C\Delta t^p \text{ (for any } \varphi \text{)}$
- "mean error" vs. "error of the mean"
- Example: for Euler-Maruyama, weak order 1, strong order 1/2 (1 when  $\sigma$  constant)

## SDEs: numerics (2)

- Trajectorial averages: estimator  $\Phi_{N_{\mathrm{iter}}} = \frac{1}{N_{\mathrm{iter}}} \sum_{n=1}^{N_{\mathrm{iter}}} \varphi(X^n)$
- Numerical scheme ergodic for the probability measure  $\psi_{\infty,\Delta t}$
- Two types of errors to compute averages w.r.t. invariant measure
   Statistical error, quantified using a Central Limit Theorem

$$\Phi_{N_{\text{iter}}} = \int \varphi \, \psi_{\infty,\Delta t} + \frac{\sigma_{\Delta t,\varphi}}{\sqrt{N_{\text{iter}}}} \, \mathscr{G}_{N_{\text{iter}}}, \qquad \mathscr{G}_{N_{\text{iter}}} \sim \mathcal{N}(0,1)$$

- Systematic errors
  - $\bullet\,$  perfect sampling bias, related to the finiteness of  $\Delta t$

$$\left|\int_{\mathcal{X}}\varphi\,\psi_{\infty,\Delta t}-\int_{\mathcal{X}}\varphi\,\psi_{\infty}\right|\leqslant C_{\varphi}\,\Delta t^{p}$$

• finite sampling bias, related to the finiteness of  $N_{
m iter}$ 

Expression of the asymptotic variance: correlations matter!

$$\sigma_{\Delta t,\varphi}^2 = \operatorname{Var}(\varphi) + 2\sum_{n=1}^{+\infty} \mathbb{E}\Big(\widetilde{\varphi}(X^n)\widetilde{\varphi}(X^0)\Big), \qquad \widetilde{\varphi} = \varphi - \int \varphi \,\psi_{\infty,\Delta t}$$

where 
$$\operatorname{Var}(\varphi) = \int_{\mathcal{X}} \widetilde{\varphi}^2 \psi_{\infty,\Delta t} = \int_{\mathcal{X}} \varphi^2 \psi_{\infty,\Delta t} - \left( \int_{\mathcal{X}} \varphi \psi_{\infty,\Delta t} \right)^2$$
  
• Note also that  $\sigma_{\Delta t,\varphi}^2 \sim \frac{2}{\Delta t} \mathbb{E} \left[ \int_0^{+\infty} \widetilde{\varphi}(X_t) \widetilde{\varphi}(X_0) \, dt \right]$ 

• Estimation with block averaging for instance, or approximation of integrated autocorrelation

## Langevin-like dynamics

## Overdamped Langevin dynamics

• SDE on the configurational part only (momenta trivial to sample)

$$dq_t = -\nabla V(q_t) \, dt + \sqrt{\frac{2}{\beta}} dW_t$$

 $\bullet$  Invariance of the canonical measure  $\nu(dq)=\psi_0(q)\,dq$ 

$$\psi_0(q) = Z^{-1} e^{-\beta V(q)}, \qquad Z = \int_{\mathcal{D}} e^{-\beta V(q)} dq$$

- Generator  $\mathcal{L} = -\nabla V(q) \cdot \nabla_q + \frac{1}{\beta} \Delta_q$ 
  - invariance of  $\psi_0$ : adjoint  $\mathcal{L}^* \varphi = \operatorname{div}_q \left( (\nabla V) \varphi + \frac{1}{\beta} \nabla_q \varphi \right)$
  - elliptic generator hence irreducibility and ergodicity
- Discretization  $q^{n+1} = q^n \Delta t \nabla V(q^n) + \sqrt{\frac{2\Delta t}{\beta}} G^n$  (+ Metropolization)

## Langevin dynamics (1)

• Stochastic perturbation of the Hamiltonian dynamics

$$\begin{cases} dq_t = M^{-1} p_t \, dt \\ dp_t = -\nabla V(q_t) \, dt - \gamma M^{-1} p_t \, dt + \sigma \, dW_t \end{cases}$$

- $\gamma, \sigma$  may be matrices, and may depend on q
- Generator  $\mathcal{L} = \mathcal{L}_{\mathrm{ham}} + \mathcal{L}_{\mathrm{thm}}$

$$\mathcal{L}_{\text{ham}} = p^T M^{-1} \nabla_q - \nabla V(q)^T \nabla_p = \sum_{i=1}^{dN} \frac{p_i}{m_i} \partial_{q_i} - \partial_{q_i} V(q) \partial_{p_i}$$
$$\mathcal{L}_{\text{thm}} = -p^T M^{-1} \gamma^T \nabla_p + \frac{1}{2} \left(\sigma \sigma^T\right) : \nabla_p^2 \qquad \left( = \frac{\sigma^2}{2} \Delta_p \text{ for scalar } \sigma \right)$$

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• Irreducibility can be proved (control argument)

## Langevin dynamics (2)

• Invariance of the canonical measure to conclude to ergodicity?

Fluctuation/dissipation relation

$$\sigma \sigma^T = \frac{2}{\beta} \gamma$$
 implies  $\mathcal{L}^* \left( e^{-\beta H} \right) = 0$ 

• Proof for scalar  $\gamma, \sigma$ : a simple computation shows that

$$\mathcal{L}_{\text{ham}}^* = -\mathcal{L}_{\text{ham}}, \qquad \mathcal{L}_{\text{ham}}H = 0$$

• Overdamped Langevin analogy  $\mathcal{L}_{thm} = \gamma \left( -p^T M^{-1} \nabla_p + \frac{1}{\beta} \Delta_p \right)$ 

 $\rightarrow$  Replace q by p and  $\nabla V(q)$  by  $M^{-1}p$ 

$$\mathcal{L}_{\text{thm}}^* \left[ \exp\left( -\beta \frac{p^T M^{-1} p}{2} \right) \right] = 0$$

• Conclusion:  $\mathcal{L}^*_{\text{ham}}$  and  $\mathcal{L}^*_{\text{thm}}$  both preserve  $e^{-\beta H(q,p)} dq dp$ Gabriel Stoltz (ENPC/INRIA)

## Langevin dynamics (3)

- Exponential convergence of semigroup  $e^{t\mathcal{L}}$  on Banach spaces  $E \cap L^2_0(\mu)$ 
  - Lyapunov techniques<sup>1</sup> on  $L_W^{\infty}(\mathcal{E}) = \left\{ \varphi \text{ measurable}, \left\| \frac{\varphi}{W} \right\|_{L^{\infty}} < +\infty \right\}$
  - Hypocoercive^2 setup  $H^1(\mu),$  with hypoelliptic regularization 3, or directly 4  $L^2(\mu)$
  - Coupling techniques<sup>5</sup>
- Allows to define the asymptotic variance (with  $\Pi \varphi = \varphi \mathbb{E}_{\mu}(\varphi)$ )

$$\sigma_{\varphi}^{2} = 2 \int_{0}^{+\infty} \int \left( e^{t\mathcal{L}} \Pi \varphi \right) \Pi \varphi \, d\mu \, dt = 2 \int (-\mathcal{L}^{-1} \Pi \varphi) \Pi \varphi \, d\mu$$

<sup>1</sup>L. Rey-Bellet, *Lecture Notes in Mathematics* (2006), Hairer/Mattingly (2011)
<sup>2</sup>Villani (2009) and before Talay (2002), Eckmann/Hairer (2003), Hérau/Nier (2004)
<sup>3</sup>F. Hérau, *J. Funct. Anal.* 244(1), 95-118 (2007)
<sup>4</sup>Dolbeault, Mouhot and Schmeiser (2009, 2015); Armstrong and Mourrat (2019)
<sup>5</sup>Eberle, Guillin and Zimmer (2019)

## Numerical integration of the Langevin dynamics (1)

• Splitting strategy: Hamiltonian part + fluctuation/dissipation

$$\begin{cases} dq_t = M^{-1} p_t dt \\ dp_t = -\nabla V(q_t) dt \end{cases} \oplus \begin{cases} dq_t = 0 \\ dp_t = -\gamma M^{-1} p_t dt + \sqrt{\frac{2\gamma}{\beta}} dW_t \end{cases}$$

- Hamiltonian part integrated using a Verlet scheme
- Analytical integration of the fluctuation/dissipation part

$$d\left(\mathrm{e}^{\gamma M^{-1}t}p_t\right) = \mathrm{e}^{\gamma M^{-1}t}\left(dp_t + \gamma M^{-1}p_t\,dt\right) = \sqrt{\frac{2\gamma}{\beta}}\mathrm{e}^{\gamma M^{-1}t}\,dW_t$$

so that

$$p_t = e^{-\gamma M^{-1}t} p_0 + \sqrt{\frac{2\gamma}{\beta}} \int_0^t e^{-\gamma M^{-1}(t-s)} dW_s$$
  
It can be shown that  $\int_0^t f(s) dW_s \sim \mathcal{N}\left(0, \int_0^t f(s)^2 ds\right)$ 

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## Numerical integration of the Langevin dynamics (2)

• Trotter splitting (define  $\alpha_{\Delta t} = e^{-\gamma M^{-1} \Delta t}$ , choose  $\gamma M^{-1} \Delta t \sim 0.01 - 1$ )

$$\begin{cases} p^{n+1/2} = p^n - \frac{\Delta t}{2} \nabla V(q^n), \\ q^{n+1} = q^n + \Delta t \, M^{-1} p^{n+1/2}, \\ \tilde{p}^{n+1} = p^{n+1/2} - \frac{\Delta t}{2} \nabla V(q^{n+1}), \\ p^{n+1} = \alpha_{\Delta t} \tilde{p}^{n+1} + \sqrt{\frac{1 - \alpha_{2\Delta t}}{\beta}} M \, G^n, \end{cases}$$

Error estimate on the invariant measure  $\mu_{\Delta t}$  of the numerical scheme There exist a function f such that, for any smooth observable  $\psi$ ,  $\int_{\mathcal{E}} \psi \, d\mu_{\Delta t} = \int_{\mathcal{E}} \psi \, d\mu + \Delta t^2 \int_{\mathcal{E}} \psi \, f \, d\mu + \mathcal{O}(\Delta t^3)$ 

• Strang splitting more expensive and not more accurate

### Metastability: large variances...



Need for variance reduction techniques!

## Variance reduction

## Main strategies for variance reduction

- **Example:** computation of the integral  $\int_{[-1/2, 1/2]^d} f$ 
  - Estimation with i.i.d. variables  $X^i \sim \mathcal{U}([-1/2, 1/2]^d)$  as  $S_{N_{\text{iter}}} = N_{\text{iter}}^{-1} \left( f(X^1) + \dots + f(X_{N_{\text{iter}}}) \right)$
  - Asymptotic variance  $\sigma_f^2 = \operatorname{Var}(f) \to \text{reduce it}?$
- Various methods (i.i.d. context, but can be extended to MCMC)
  - Antithetic variables  $I_{N_{\text{iter}}} = \frac{1}{2N_{\text{iter}}} \sum_{i=1}^{N_{\text{iter}}} \left( f\left(X^{i}\right) + f\left(-X^{i}\right) \right)$
  - $\bullet$  Control variates with  $\sigma_{f-g}^2 \ll \sigma_f^2$  and g analytically integrable

$$I_{N_{\text{iter}}} = \frac{1}{N_{\text{iter}}} \sum_{i=1}^{N_{\text{iter}}} (f - g) \left( X^{i} \right) + \int_{[-1/2, 1/2]^{d}} g$$

Stratification: partition domain, sample subdomains, aggregateImportance sampling

## Importance sampling

- Importance sampling function  $\widetilde{V}$ 
  - Target measure  $\pi_0(dx) = Z_0^{-1} \mathrm{e}^{-V(x)} \, dx$
  - $\bullet$  Sample a modified target measure  $\pi_{\widetilde{V}}(dx)=Z_{\widetilde{V}}^{-1}{\rm e}^{-(V+\widetilde{V})(x)}\,dx$
  - Reweight sample points  $x^n \sim \pi_{\widetilde{V}}$  by  $e^{\widetilde{V}}$

$$\widehat{\varphi}_{N_{\text{iter}},\widetilde{V}} = \frac{\sum_{n=1}^{N_{\text{iter}}} \varphi(x^n) e^{\widetilde{V}(x^n)}}{\sum_{n=1}^{N_{\text{iter}}} e^{\widetilde{V}(x^n)}} \xrightarrow[N_{\text{iter}} \rightarrow +\infty]{\text{a.s.}} \frac{\int \varphi e^{\widetilde{V}} d\pi_{\widetilde{V}}}{\int e^{\widetilde{V}} d\pi_{\widetilde{V}}} = \int \varphi d\pi_0$$

- In practice, replace  $-\nabla V$  with  $-\nabla V \nabla \widetilde{V}$  (in Langevin, MALA, etc)
- A good choice of the importance sampling function can improve the performance of the estimator... but a bad choice can degrade it!

## High dimensional importance sampling

#### • General strategy:

- find some low-dimensional (nonlinear) function  $\xi(x)$  which encodes the metastability of the sampling method
- $\bullet$  bias by the associated free energy:  $\widetilde{V}(x)=F(\xi(x))$  with

$$e^{-F(z)} = \int e^{-V(x)} \delta_{\xi(x)-z}(dx)$$

• Simple case:  $\xi(x) = x_1$ , in which case

$$F(z) = -\ln\left(\int e^{-V(z,x_2,\dots,x_d)} dx_2\dots dx_d\right)$$

• Various methods to compute the free energy: thermodynamic integration, umbrella sampling, adaptive methods, ...

### Free energy biasing for Bayesian inference



[CLS12] N. Chopin, T. Lelièvre and G. Stoltz, Statist. Comput., 2012